DIFFERENT MODELS FOR ESTIMATING GENETIC PARAMETERS FOR GROWTH TRAITS OF HOLSTEIN-FRIESIAN UNDER SEMITROPICAL CONDITIONS

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SUMMARY

Estimates of genetic parameters resulting from three different models for birth weight, weaning weight, yearling weight and total gain were compared,

Data consisted of 823 Holstein-Friesian animals progenies of 25 sires and 459 dams in a single herd. The MTDFREML programs were used to estimate heritabilities, predicted transmitting abilities (PTA) and genetic correlations. The three models compared were: Model I, a sire model, included year-season of birth and sex of calf as fixed effects and sire as random effect. Model II, a sire-dam model, included, in addition to the effects in model I, dam as a second random effect. Model III, a full animal model, included, in addition to the fixed effects of model I, animal additive direct genetic effect, maternal genetic effect and permanent environmental effect, as three random effects.

The likelihood function showed that the full animal model best fit the data. Heritability estimates and genetic correlations obtained from model III were higher than those obtained from models I and II for all traits. Models I and II resulted in changes in ranking for PTA in contrast with model III.

In conclusion, full animal model could be most appropriate for estimating genetic parameters for growth traits than the sire or the sire-dam models.

Keywords: Growth, animal model, heritability, genetic correlation, predicted transmitting ability

INTRODUCTION

Although Holstein-Friesian is considered as dairy animals, the growth of these animals is of great importance. Extra male calves are grown as meat animals and female calves are grown as replacement heifers. Studies have indicated that when prepubertal growth rates of heifers increase, time to conception, age at first calving, and milk yield during first lactation decrease (Swanson, 1960, Grander *et al.*, 1977, and Little and Kay, 1979). During the prepubertal period of heifer growth, the mammary gland is sensitive to body weight gain (Sejrsen *et al.*, 1982, Sejrsen *et al.*, 1983, Peticlerc *et al.*, 1984). Koenen and Groen (1996) reported that the expected increase in profitability was 2 to 4% when body weight was included in the breeding

goal and in the selection index due to the relationship between body weight and production traits. Van Amburgh *et al.* (1998) indicated that milk yield was significantly reduced for heifers grown at 0.94 kg/d compared with that of heifers grown at 0.68 kg/d. However, Mercadante *et al.* (2003) concluded that selection for body weight promoted high and consistent weight response for both yearling and later ages, without compromising the reproductive performance of the cows with respect to days to calving and calving success. There is evidence that changes in performance of animals with age are influenced by genetic factors (Atchley *et al.*, 1997 and Atchley, 1998). Growth traits in cattle are important in selection program. Therefore, estimating the genetic parameters for some growth traits and implement them in a selection program would be of great importance.

The estimation of variance and covariance components, hence estimating genetic parameters, evolved from Henderson's Methods I and III (Henderson, 1953), through the mixed models methodology (Henderson, 1988). The Restricted Maximum Likelihood, REML, (co)variance component estimation is the most commonly used algorithm in such estimation utilizing different mixed models.

The objectives of the present work were to compare estimates of genetic parameters for birth weight (BWT), weaning weight (WWT), yearling weight (YWT) and total gain (TG) from Holstein-Friesian cattle using different statistical models to determine whether simpler models produce estimates similar to those produced by more complex alternatives.

MATERIAL AND METHODS

Data were 823 records for Holstein-Friesian cattle from one herd owned by the Egyptian Company for Meat and Milk Production, Faraskour, Damietta collected from 1988 to 1994. Animals were housed in semi-roofed yards. Animals were feed after weaning on concentrate mixture and roughage {such as rice straw and green Egyptian clover (*Trifolium alexandrinum*) or its hay}. Feeding allowances were calculated according to NRC (1981). Fresh and clean water was available all times.

Data included 271 male calves and 552 female calves. These calves were progenies of 25 sires and 459 dams. Two seasons were defined, from April to September and from October to March. Hence, 14 year-seasons were identified. Four traits were selected for analyses: birth weight (BWT), weaning weight (WWT) at 100 days of age, yearling weight (YWT) and total gain (TG) from birth to 12 months of age. Table 1 lists descriptive statistics of the records for the four traits evaluated for whole data and for both sexes.

MODELS:

Three models were used to estimate genetic and phenotypic parameters for the four traits. They included heritability for each trait and genetic and phenotypic correlations between different traits. The first model (Model I) was the sire model. The second (Model II) was the sire-dam model. The third (Model III) was the animal model that included the animal direct genetic effect, the maternal genetic effect, and the permanent environmental effect.

The equation for model III for multiple traits was as follows:

$$y = Xb + Za + Mm + Wp + e$$

where:

y is a vector of observations of the four traits:

b is a vector of fixed effects (year-season of birth and sex of calf):

a is a vector of random additive direct genetic effects;

m and p are vectors of random maternal genetic and permanent environmental effects;

e is a vector of random residual effect; and

X, Z. M and W are known incidence matrices relating observations to the respective fixed and random effects with Z and M augmented with columns of zeros for animals without records.

The first and the second moments of the model were assumed to be:

$$E(y) = Xb$$
, $E(a) = E(m) = E(p) = E(e) = 0$; and

$$V\begin{bmatrix} a_{I} \\ m_{J} \\ p_{I} \\ e_{I} \end{bmatrix} = \begin{bmatrix} A \otimes J\sigma_{u_{I}} & 0 & 0 & 0 \\ 0 & A \otimes J\sigma_{m_{I}} & 0 & 0 \\ 0 & 0 & I_{I_{I}} \otimes \sigma_{P_{I}}^{2} & 0 \\ 0 & 0 & 0 & I_{I_{I}} \otimes \sigma_{e_{I}}^{2} \end{bmatrix}$$

A is the additive numerator relationship matrix for animals. J is a matrix of I's of order 4x4. I_{nc} and I_n are identity matrices of order equal to the number of dams (nc) and to the total number of animals (n), respectively. \bigotimes means the direct product of two matrices. Subscripts i and j refer to the four traits (i = j = 1 to 4). When i = j, it means the diagonal which is the variance. When $i \neq j$, it means the off diagonal, which is the covariance. All random effects were considered uncorrelated.

Model I (sire model) and model II (sire-dam model) included the same fixed effects as model III. Model I included the sire only as random effect. Model II included the sire and the dam as two random effects. Therefore, the differences among the models refer to the number of random effects considered.

Genetic parameters were estimated for the three models with derivative-free REML (Meyer, 1989) using multiple-trait derivative free restricted maximum likelihood (MTDFREML) programs (Boldman *et al.* 1995). The method involves maximizing the likelihood function (A) given the data and is the same as maximizing log Λ or minimizing -2 log Λ . Heritability estimates obtained from model II were calculated based on sire and dam variance components as $2(\sigma_s^2 + \sigma_d^2)/\sigma_p^2$, where σ_s^2 ,

 σ_d^2 , and σ_p^2 were sire, dam and total variance components, respectively.

To compare models each trait was reanalyzed as a single trait with the three different models to obtain the log likelihood value. It was assumed that the higher the likelihood function the more the model explained the data. Likelihood function is higher when new parameters are added to the model. Comparisons of the different models were made with likelihood-ratio test. The differences between the function values for pairs of models can be tested against the chi-square distribution with degrees of freedom being the difference in number of variance or covariance components in the models (Dobson, 1990). This method is based on the property that the difference $-2[\log A_i - \log A_{i'}]$ has chi-squared distribution, where A_i and $A_{i'}$ are the values of likelihood function.

Table 1. Descriptive statistics for records for birth weight (BWT), wearing weight (WWT), yearling weight (YWT) and total gain (TG), kg

| Item | BWT | WWT | YWT | TG |
|------------------------------|------|------|-------|-------|
| Whole data (823 records): | | | | |
| Mean | 30.3 | 74.4 | 219.4 | 189.1 |
| Standard deviation | 4.9 | 13.5 | 32.8 | 31.5 |
| Coefficient of variation % | 16.1 | 18.1 | 14.9 | 16.7 |
| Male calves (271 records): | | | | |
| Mean | 32 | 76.3 | 223.5 | 191.5 |
| Standard deviation | 4.8 | 14 | 34 | 33 |
| Coefficient of variation % | 15 | 18.3 | 15.2 | 17.2 |
| Female calves (552 records): | | | | |
| Mean | 29.4 | 73.4 | 217.3 | 187,9 |
| Standard deviation | 4.7 | 13.1 | 32 | 30.7 |
| Coefficient of variation % | 15.8 | 17.9 | 14.7 | 16.3 |

RESULTS AND DISCUSSION

The definition of the correct model that better fits the data is very important. Table 2 shows the calculated chi-square values for the difference between different models in the log likelihood functions. This test showed a significant difference between the animal model and both the sire and sire-dam models for all traits except birth weight and weaning weight where there was no significance difference between the animal model and the sire-dam model for both traits. The log likelihood showed a better fit with the animal model than with sire or sire-dam models. These results agree with Ferreira *et al.* (1999).

Table 3 shows estimates of heritability and genetic correlation from different models. In general, there was an increase in the estimate of heritability for all traits by adding more random effects to the model with different magnitude. They ranged from 0.27 and 0.49 for birth weight, 0.17 and 0.28 for weaning weight, 0.14 and 0.32 for yearling weight and 0.11 and 0.21 for total gain, using the sire model and the animal model, respectively. Sire-dam model heritability estimates were within these ranges for all traits. Estimates of heritability from the animal model were comparable to that reported by Meyer (1993) and Koots *et al.* (1994).

Table 2. Calculated chi-square values for likelihood test used to compare different models (sire model, Model I, sire dam mode, Model II and animal model, Model III) for birth weight (BWT), weaning weight (WWT), yearling weight (YWT), and total gain (TG)

| | Comparison between models | | | | |
|-------|---------------------------|--------|--------------------|--|--|
| Trait | | | | | |
| BWT | 75 | 76.76 | 1.76 ^{NS} | | |
| WWT | 115 | 118,25 | 3.25 ^{NS} | | |
| YW1 | 210 | 404 | 194*** | | |
| TG | 195* | 346 | 151 | | |

Significant difference at p < 0.01

Table 3. Heritability (diagonal) estimates and genetic correlations (below diagonal) among birth weight (BWT), weaning weight (WWT), yearling weight (YWT), and total gain (TG) using different models

| Trait | BWT | WWT | YWT | TG |
|-------|-------|----------|------------------|------|
| | ··· | Mode | I (Sire model) | |
| BWT | 0.27 | | | |
| WWT | 0.46 | 0.17 | | |
| YWT | 0.22 | 0.39 | 0.14 | |
| TG | 0.39 | 0.31 | 0.79 | 0.11 |
| | | Model II | (Sire-Dam model) | |
| BWT | 0.29 | | | |
| WWT | 0.51 | 0.18 | | |
| YWT | 0.24 | 0.41 | 0.21 | |
| TG | 0.41 | 0.38 | 0.88 | 0.14 |
| | | Model II | I (Animal model) | |
| BWT | ().49 | | | |
| WWT | 0.65 | 0.28 | | |
| YWT | 0.31 | 0.43 | 0.32 | |
| TG | 0.45 | 0.48 | 0.95 | 0.21 |

Genetic correlation showed same trend as the heritability estimates. Genetic correlation estimates obtained from the animal model were in general higher than that obtained from the sire or the sire-dam model. It ranged from 0.22 between birth weight and yearling weigh using the sire model to 0.95 between yearling weight and total gain using the animal model.

Table 4 shows the product moment and rank correlations between PTA values using the animal model and with the sire and the sire-dam models. Correlation between PTA values and ranks using the animal model with those using the sire-dam model were higher than the correlations with the sire model for all traits. These results were similar to that obtained by Ferreira et al. (1999). They suggest that the sire and/or the sire-dam models may not be adequate for estimating genetic parameters for growth traits if the full animal model can be assumed to be most appropriate.

Not significant at p > 0.05

Table 4. Product moment and rank correlations between PTA values using the animal model (Model II) and using the sire-model (Model I) and the sire-dammodel (Model II) for birth weight, weaning weight, yearling weight and total gain

| - S | | |
|----------------------------|-----------------------|------------------------|
| | Model III and Model I | Model III and Model II |
| | | |
| Birth weight: | | |
| Product moment correlation | 0.76 | 0.82 |
| Rank correlation | 0.74 | 0.81 |
| Weaning weight: | | |
| Product moment correlation | 0.67 | 0.73 |
| Rank correlation | 0.62 | 0.69 |
| Yearling weight: | | |
| Product moment correlation | 0.58 | 0.63 |
| Rank correlation | 0.55 | 0.61 |
| Total gain: | | |
| Product moment correlation | 0.64 | 0.68 |
| Rank correlation | 0.61 | 0.63 |

CONCLUSIONS

Selection of the correct model is very important, because the more complex the model, the larger the time needed for solution. However, increasing the accuracy of estimation by adding more effects to the model that better explain the data would be more important than saving the time needed for solution. Full animal model generally had higher estimates of heritability and genetic correlation for birth weight, weaning weight, yearling weight and total gain than sire or sire-dam models. In addition, the likelihood function test showed a better fit with the full animal model than the sire or the sire-dam models. However, there was no significance difference in the likelihood function between the animal model and the sire-dam model for birth weight and weaning weight. Correlations and ranks between estimated TA values from the animal model and the sire and the sire-dam models suggest that that the animal model could be more appropriate for estimating genetic parameters for growth traits.

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نماذج مختلفة لتقدير المعالم الوراثية لصفات النمو للهولستين – فرزيان تحت الظروف شبه الحارة

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نمت مقارنة المعالم الوراثية المقدرة الناتجة من ثلاثة نماذج مختلفة لكل من وزن الميلاد، وزن الفطــــام، الوزن عند عمر سنة، وإجمالي الزيادة الوزنية.

استخدمت بيانات ٨٢٣ حيوان نسل ٢٥ طلوقة و ٤٥٩ أم فى قطيع واحد. استخدم فى التحليل طريقسسة MTDFRML لحساب المكافئات الوراثية والقيم المتوقعة للقدرة المرورية لكل صفة وكذلك الارتبساط الوراثي بين الصفات المختلفة.

تمت مقارنة ثلاثة نماذج للتحليل: النموذج الأول (نموذج الطلوقة) اشتمل على سنة -موسم ميلاد الحيوان ، جنس الحيوان كمؤثرين ثابتين والطلوقة كمؤثر عشوائي. النموذج الثاني (نموذج الطلوقسة -الأم) تضمن المؤثرات المذكورة في النموذج الأول بالإضافة إلى الأم كمؤثر عشوائي ثاني. النموذج الشالث (نموذج الحيوان) اشتمل على المؤثرات الثابتة المذكورة في النموذج الأول بالإضافة إلى ثلاثة مؤثرات عشوائية هي الأثر الوراثي الأمي ، الأثر البيئ الدائم.

أظهرت الدالة الاحتمالية المستونة likelihood function أن نموذج الحيوان كان الأفضل في تمثيل البيانات. تقديرات المكافئات الوراثية والارتباط الوراثي المحسوبة من نموذج الحيوان كانت أعلى من مثيلتها المحسوبة من كل من نموذج الطلوقة ونموذج الطلوقة - الأم. أظهر النموذجين الأول والثاني اختلافات في ترتبب القدرة المرورية المنوقعة مقارنة بالنموذج الثالث.

خلصت الدراسة إلى أن نموذج الحيوان هو الأفضل لتقييم الطلائق لصفات النمو مقارنة بنموذج الطلوقة أو نموذج الطلوقة و الأم.